

# On a Conjecture Concerning a Theorem of Cramér and Wold\*

Guenther Walther

*Stanford University*

A conjecture concerning the Cramér–Wold device is answered in the negative by giving a Fourier-free, probabilistic proof using only elementary techniques. It is also shown how a geometric idea allows one to interpret the Cramér–Wold device as a special case of a more general concept. © 1997 Academic Press

## 1. INTRODUCTION

A fundamental and widely used theorem states that for checking convergence in distribution of multivariate random variables it is enough to check convergence for all one-dimensional projections. More specifically, the so-called “Cramér–Wold device,” due to Cramér and Wold [3] where the technique was initiated, establishes the following two assertions:

(I) A probability measure on Euclidean space is uniquely determined by the values it gives to halfspaces.

(II) In Euclidean  $d$ -space, a sequence of random variables  $X_n$  converges in distribution to a random variable  $X$  if and only if  $\langle a, X_n \rangle$  converges in distribution to  $\langle a, X \rangle$  for each  $a \in \mathbf{R}^d$ .

Both theorems, (I) and the stronger (II), although they are very simple in their statements, have been conjectured to require Fourier analysis for their proofs; see, e.g., p. 396 of Billingsley [2] for the first and p. 49 of Billingsley [1] for the second part.

This note gives probabilistic proofs of the two theorems and thus answers this conjecture to the negative. The main argument of the proof is a simple probabilistic idea that goes back to the early stages of probability theory. Also, a geometric idea that belongs to the standard repertoire in

Received February 19, 1997.

AMS 1991 subject classification numbers: 60B10, 60E10.

Key words and phrases: Cramér–Wold device, determining class, Fourier analysis, halfspace, polar set.

\* This research was partially supported by NSF Grant DMS-92-24868 at U.C. Berkeley.

convex geometry shows how the Cramér–Wold device can be interpreted as a special case of a more general concept.

## 2. SOME FACTS ABOUT DETERMINING CLASSES

The setting throughout will be the Euclidean  $d$ -space  $\mathbf{R}^d$  equipped with the standard inner product  $\langle \cdot, \cdot \rangle$  and the Euclidean norm  $|\cdot|$ . Write  $\mathcal{M}^d$  for the set of probability measures on  $\mathbf{R}^d$  and  $\mathbf{bM}^d$  for the set of bounded, measurable, and real-valued functions on  $\mathbf{R}^d$ . Recall that a set  $\mathcal{D} \subset \mathbf{bM}^d$  is called a determining class if  $P, Q \in \mathcal{M}^d$  and  $\int f dP = \int f dQ$  for all  $f \in \mathcal{D}$  imply  $P = Q$ . A basic observation now is that the following lemma has a simple probabilistic proof that does not require Fourier analysis:

LEMMA 1. *Let  $f \in \mathbf{bM}^d$  be nonnegative and satisfy  $0 < \int f(x) dx < \infty$ . Then  $\{f((a - \cdot)/b), a \in \mathbf{R}^d, b > 0\}$  is a determining class.*

*Proof.* Let  $P, Q \in \mathcal{M}^d$  and assume

$$\int f\left(\frac{a-x}{b}\right) P(dx) = \int f\left(\frac{a-x}{b}\right) Q(dx) \quad \text{for all } a \in \mathbf{R}^d, b > 0. \quad (1)$$

Normalize  $f$  so that  $\int f(x) dx = 1$  and define  $F \in \mathcal{M}^d$  via its density  $f$ . Let  $X_P, X_Q$ , and  $X_F$  be independent random variables in  $\mathbf{R}^d$  with distribution  $P, Q$ , and  $F$ , respectively. Then  $X_P + \varepsilon X_F$  has density  $\varepsilon^{-d} \int f((\cdot - x)/\varepsilon) P(dx)$ , so (1) shows that  $\mathcal{L}(X_P + \varepsilon X_F) = \mathcal{L}(X_Q + \varepsilon X_F)$  for all  $\varepsilon > 0$ . Now let  $h$  be any continuous function in  $\mathbf{bM}^d$ . Then bounded convergence gives

$$\mathbb{E}h(X_P) = \lim_{\varepsilon \downarrow 0} \mathbb{E}h(X_P + \varepsilon X_F) = \lim_{\varepsilon \downarrow 0} \mathbb{E}h(X_Q + \varepsilon X_F) = \mathbb{E}h(X_Q).$$

$P = Q$  follows. ■

The idea of determining a probability measure by its convolutions with an appropriate class of measures goes back at least to Liapounoff [4] and Lindeberg [5], who employed convolutions in their proofs of the Central Limit Theorem to make use of the resulting smoothness properties.

As an aside, note that Lemma 1 can be sharpened with the use of Fourier analysis and the additional assumption that  $\int e^{i\langle t, x \rangle} f(x) dx \neq 0$  for all  $t$ : Requiring (1) only for  $b = 1$  gives  $P * F = Q * F$ . The characteristic functions of these convolutions factor, so dividing by the nonzero characteristic function of  $F$  and using the uniqueness theorem of characteristic functions shows  $P = Q$ . The resulting determining class apparently is much

smaller than the one required in Lemma 1, but in the case of interest here this is only seemingly so:

For fixed  $u \neq 0$  the function  $f(x) := 1(\langle x, u \rangle \leq 1)$  is the indicator of a halfspace and one readily checks that  $\{f(a - \cdot), a \in \mathbf{R}^d\} = \{f((a - \cdot)/b), a \in \mathbf{R}^d, b > 0\}$ , so nothing is lost by forgoing Fourier analysis in this context.

Of course the above function  $f$  is not integrable, so Lemma 1 does not apply. But an application of Fubini's theorem to the result of Lemma 1 gives

LEMMA 2. *Let  $f(x, u) \in \mathbf{bM}^{d+p}$ ,  $\mu_1, \dots, \mu_m \in \mathcal{M}^p$ , and  $a_1, \dots, a_m \in \mathbf{R}$  such that  $F(x) := \sum_{k=1}^m a_k \int f(x, u) \mu_k(du)$  is nonnegative and satisfies  $0 < \int F(x) dx < \infty$ . Then  $\{f((a - \cdot)/b, u), a \in \mathbf{R}^d, u \in \mathbf{R}^p, b > 0\}$  is a determining class.*

PROOF OF (I)

Set  $f(x, u) := 1(\langle x, u \rangle \leq 1)$ . Then Lemma 2 leads one to consider functions of the form

$$f_\mu^*(\cdot) := \int_{\langle \cdot, u \rangle \leq 1} \mu(du). \tag{2}$$

Denote by  $\Phi$  and  $\phi$  the distribution function and the density function, respectively, of the standard normal distribution on  $\mathbf{R}$ , and set  $\Phi_\sigma(\cdot) = \Phi(\cdot/\sigma)$ . We will show in a moment:

(L) There exists a linear combination  $g(t) := \sum_{k=1}^{d+1} a_k \Phi_{\sigma_k}(t) + a_{d+2}$  with

- $g(0) = 0$ ,
- $g(t)$  is strictly increasing for  $t \geq 0$ ,
- $g(t) = O(t^{d+1})$  as  $t \downarrow 0$ .

Now apply Lemma 2 with  $F(x) := \sum_{k=1}^{d+2} a_k f_{\mu_k}^*(x)$ , where  $\mu_k = N(0, \sigma_k^2 I_d)$  for  $1 \leq k \leq d+1$  and  $\mu_{d+2} = \delta_0$ . Here  $I_d$  and  $\delta_0$  denote the  $d \times d$  identity matrix and point mass at 0, respectively. Observe that  $f_{\delta_0}^* \equiv 1$ , and as projections of standard normal distributions are standard normal (which can be shown without Fourier analysis),

$$f_{N(0, \sigma^2 I_d)}^*(x) = \Phi_\sigma(1/|x|) \quad (\text{if } x = 0 \text{ interpret } \Phi_\sigma(1/|x|) = 1).$$

Hence  $F(x) = g(1/|x|) = O(1/|x|^{d+1})$  as  $|x| \rightarrow \infty$ . Together with the properties of  $g$  one sees that  $F$  is nonnegative and satisfies  $0 < \int F(x) dx < \infty$ .

The Cramér–Wold theorem (I) now follows from Lemma 2 and the fact that  $f((a-\cdot)/b, u)$  is the indicator of a closed halfspace or  $\mathbf{R}^d$  for all  $a, u \in \mathbf{R}^d$ ,  $b > 0$ .

It remains to prove (L). We will choose the  $a_k$  in the linear combination  $g(t)$  to eliminate the  $d$  coefficients of the  $t^n$ ,  $n = 1, \dots, d$ , in the Taylor series expansion about 0,

$$\Phi_\sigma(t) = 1/2 + \sum_{n=1}^d \frac{1}{n! \sigma^n} \Phi^{(n)}(0) t^n + O(t^{d+1}). \quad (3)$$

For simplicity of exposition we will not make use of the fact that  $\Phi^{(2n)}(0) = 0$  for  $n \geq 1$ . Using pairwise different  $\sigma_k > 0$  and setting  $x := \sigma_{d+1}^{-1}$  in the polynomial interpolation formula

$$x^n = \sum_{k=1}^d \sigma_k^{-n} \prod_{i=1, i \neq k}^d \left( \frac{x - \sigma_i^{-1}}{\sigma_k^{-1} - \sigma_i^{-1}} \right), \quad n = 0, \dots, d-1,$$

one sees that  $b_k := \prod_{i=1, i \neq k}^d (\sigma_{d+1}^{-1} - \sigma_i^{-1}) / (\sigma_k^{-1} - \sigma_i^{-1})$ ,  $k = 1, \dots, d$ , and  $b_{d+1} = -1$  solve the system of  $d$  equations

$$\sum_{k=1}^{d+1} \sigma_k^{-n} b_k = 0, \quad n = 0, \dots, d-1. \quad (4)$$

Employing the increasing sequence  $\sigma_k := 4^k$ , one concludes that

$$a_k := -\sigma_k b_k = (-1)^{d+1-k} 4^k \prod_{i=1, i \neq k}^d \frac{|4^{-(d+1)} - 4^{-i}|}{|4^{-k} - 4^{-i}|}, \quad k = 1, \dots, d+1,$$

solve the system (4) for  $n = 1, \dots, d$ . Hence the expansion (3) gives

$$g(t) = \sum_{k=1}^{d+1} a_k \Phi_{\sigma_k}(t) - \sum_{k=1}^{d+1} a_k / 2 = O(t^{d+1}) \quad \text{as } t \rightarrow 0.$$

Further,

$$g'(t) = \sum_{k=1}^{d+1} \frac{a_k}{\sigma_k} \phi \left( \frac{t}{\sigma_k} \right),$$

and for  $1 \leq k \leq d$ ,

$$\begin{aligned} & \left| \frac{(a_{k+1}/\sigma_{k+1}) \phi(t/\sigma_{k+1})}{(a_k/\sigma_k) \phi(t/\sigma_k)} \right| \\ &= \frac{\prod_{i=1, i \neq k+1}^d |4^{-(d+1)} - 4^{-i}| \cdot \prod_{i=1, i \neq k}^d |4^{-(k+1)} - 4^{-(i+1)}| \cdot 4^{d-1}}{\prod_{i=1, i \neq k}^d |4^{-(d+1)} - 4^{-i}| \cdot \prod_{i=0, i \neq k}^{d-1} |4^{-(k+1)} - 4^{-(i+1)}|} \\ & \quad \times \frac{\phi(t/4\sigma_k)}{\phi(t/\sigma_k)} \\ &= \frac{|4^{-(d+1)} - 4^{-k}| \cdot 4^{d-1} \cdot (\phi(t/\sigma_k))^{1/16}}{|4^{-(k+1)} - 4^{-1}| \cdot \phi(t/\sigma_k)} \\ &= \frac{|4^{-1} - 4^{d-k}| \cdot 4^k}{|1 - 4^k|} \cdot \left( \phi \left( \frac{t}{\sigma_k} \right) \right)^{-15/16} \\ & \geq 1, \end{aligned}$$

because  $|4^{-1} - 4^{d-k}| \geq 3/4$  and  $\phi(t/\sigma_k) \leq \phi(0) = 1/\sqrt{2\pi} \leq (3/4)^{16/15}$ .

As the signs of the  $a_k$ ,  $k \geq 1$ , are alternating with the sign of  $a_{d+1}$  being positive, it follows that  $g'(t) > 0$  for  $t > 0$ . Clearly,  $g(0) = 0$ .

#### 4. PROOF OF (II) AND A GENERALIZATION

Part II of the Cramér–Wold theorem follows readily from Part I:  $X_n \xrightarrow{\text{dist}} X$  implies

$$\langle a, X_n \rangle \xrightarrow{\text{dist}} \langle a, X \rangle \quad \text{for each } a \in \mathbf{R}^d \tag{5}$$

by the continuous mapping theorem. Conversely, suppose (5) holds. Let  $\{e_1, \dots, e_d\}$  be an orthonormal system in  $\mathbf{R}^d$ . For  $\delta > 0$ ,  $\bigcap_{i=1}^d \{x: \langle \delta e_i, x \rangle \leq 1\} \cap \bigcap_{i=1}^d \{x: \langle -\delta e_i, x \rangle \leq 1\}$  is a cube centered at 0 with sidelength  $2/\delta$ . Hence a variation of Boole’s inequality together with (5) shows that the sequence  $\{\mathcal{L}(X_n)\}$  is uniformly tight. By Prohorov’s theorem and the subsequence criterion for metric spaces it is therefore enough to show that any weakly convergent subsequence  $\{\mathcal{L}(X_{n_k})\}$  converges to  $\mathcal{L}(X)$ . But this follows from the already proved implication (5) together with the uniqueness theorem (I).

There is a fundamental geometric concept involved in (2) that allows the Cramér–Wold theorem to be interpreted as a special case of a more general statement:

The *dual (polar) set* of a set  $X \in \mathbf{R}^d$  is defined as

$$X^* := \{u \in \mathbf{R}^d : \langle x, u \rangle \leq 1 \text{ for all } x \in X\},$$

see e.g. Stoer and Witzgall [6]. If  $x \in \mathbf{R}^d \setminus \{0\}$  then  $\{x^*\} = \{u \in \mathbf{R}^d : \langle x, u \rangle \leq 1\}$  is a closed halfspace containing 0 in its interior; if  $x=0$  then  $\{x^*\}$  is all of  $\mathbf{R}^d$ . This geometric concept leads one to define for a probability measure  $\mu \in \mathcal{M}^d$  the *dual measure*  $\mu^*$  via its density  $f_\mu^*$  given in (2). One checks that  $f_\mu^*$  is upper semicontinuous, hence is measurable. Thus  $f_\mu^*$  is indeed the density of a  $\sigma$ -finite measure  $\mu^*$ .  $\mu^*$  is always an infinite measure. See Walther [7], where also statistical motivations are given for constructing such measures.  $\mu^*$  can formally also be motivated as follows: For simplicity consider a one-dimensional setting and let  $F$  denote the distribution function of a probability measure. For real  $x$  write

$$\begin{aligned} F(x) &= \int 1_{(-\infty, x]}(u) F(du) \\ &= \int 1_{[u, \infty)}(x) F(du). \end{aligned} \tag{6}$$

Formally, (6) can be read as a mixture of uniform densities (albeit not of probability densities).

Now the Cramér–Wold theorem is a consequence of the following more general theorem about dual measures a proof of which can be found in Walther [8]:

**THEOREM 1.** *Let  $X, X_1, X_2, \dots, \in \mathbf{R}^d$  be a sequence of random variables with  $\mathcal{L}(X) = F$ ,  $\mathcal{L}(X_n) = F_n$ ,  $n \geq 1$ . Then the following are equivalent:*

- (i)  $F_n \xrightarrow{\text{weakly}} F$
- (ii)  $\langle a, X_n \rangle \xrightarrow{\text{dist}} \langle a, X \rangle$  for all  $a \in \mathbf{R}^d$
- (iii)  $f_{F_n}^* \xrightarrow{\text{a.e.}} f_F^*$
- (iv)  $F_n^* \rightarrow F^*$  in variation norm on compacts
- (v)  $F_n^* \xrightarrow{\text{vaguely}} F^*$ .

If  $F_n$  is the empirical measure of  $F$ , then (iii) can be strengthened to uniform convergence  $F$ -almost surely.

As a corollary one obtains the following identifiability property:  $F^* = G^*$  iff  $F = G$ .

## ACKNOWLEDGMENT

I thank Richard Olshen for his interest in the paper and for valuable suggestions.

## REFERENCES

- [1] Billingsley, P. (1968). *Convergence of Probability Measures*. Wiley, New York.
- [2] Billingsley, P. (1986). *Probability and Measure*, 2nd ed. Wiley, New York.
- [3] Cramér, H., and Wold, H. (1936). Some theorems on distribution functions. *J. London Math. Soc.* **11** 290–294.
- [4] Liapounoff, A. M. (1900). Sur une proposition de la théorie des probabilités. *Bull. Acad. Impériale Sci. St. Pétersbourg* **13** 359–386.
- [5] Lindeberg, J. W. (1922). Eine neue Herleitung des Exponentialgesetzes in der Wahrscheinlichkeitsrechnung. *Math. Z.* **15** 211–225.
- [6] Stoer, J., and Witzgall, C. (1970). *Convexity and Optimization in Finite Dimensions*, Vol. 1. Springer-Verlag, Berlin/Heidelberg.
- [7] Walther, G. (1995). Monte Carlo sampling in dual space for approximating the empirical halfspace distance. *Ann. Statist.*, in press.
- [8] Walther, G. (1994). *Statistical Applications of Geometric Duality*. Ph.D. thesis, Dept. of Statistics, University of California, Berkeley.